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## **Trade Liberalization and Endogenous Quality Choice in Food and Agricultural Trade**

**Jihyun Eum and Ian Sheldon**

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# Trade Liberalization and Endogenous Quality Choice in Food and Agricultural Trade

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# Introduction

- Trade standards exist to protect safety and environment: importance of food safety and its quality has been emphasized
- Stylized facts
  - Hidden causes and consequences of international trade at country-level can be interpreted with firm heterogeneity
  - Differences in productivity among firms causes changes in trade participation
- Literature review
  - Trade liberalization and quality sorting: Amiti and Khandelwal (2013), Fan et al. (2014)
  - Endogenous quality choice model: Baldwin and Harrigan (2011), Johnson(2012), Kugler and Verhoogen (2012)

# Research Question and Contribution

- Illustrate heterogeneous firm trade model with endogenous quality choice
- Estimate model with agricultural and food trade data
  - Evaluate the determinants of bilateral trade
  - Analyze the effect of food safety standards as a fixed trade cost
- Contribution
  - Introduce the impact of selection into exporting with consideration of product quality in agricultural and food trade
  - Advanced standards data: World Integrated Trade Solution (WITS) and Integrated Trade Intelligence Portal (I-TIP) from WTO

# Theoretical Background

- Preferences

$$U = \left[ \int_{\omega \in \Omega} (q(\omega)x(\omega))^{(\sigma-1)/\sigma} d\omega \right]^{\sigma/(\sigma-1)} \quad \text{where } \sigma > 1 \quad (1)$$

$$x(\omega) = p(\omega)^{-\sigma} q(\omega)^{(\sigma-1)} A \quad \text{where } A = EP^{(\sigma-1)} \quad (2)$$

- P is aggregated price index
- E is aggregated consumption
- Firms are heterogeneous in (1) productivity (a) and (2) product quality (q)
  - J countries,  $N_j$  firms under monopolistic competition
  - Marginal cost of production:  $\frac{c_j}{a}$
  - Firms choose optimal domestic price ( $p_j$ ) and export price ( $p_{ij}^x$ )

$$p_j = \frac{\sigma}{\sigma-1} (c_j/a) \quad \& \quad p_{ij}^x = \frac{\sigma}{\sigma-1} \frac{c_j \tau_{ij}}{a}$$

# Theoretical Background

- Productivity and quality are linked as below (Baldwin and Harrigan, 2011)

$$q = a^{\theta-1} \quad \text{where} \quad \theta - 1 > 0$$

$\theta - 1$  is “quality elasticity” or “scope for quality differentiation”

- Profit and zero-profit condition

$$\pi_{ij}(a) = \frac{1}{\sigma} \left( \frac{\sigma}{\sigma - 1} \right)^{1-\sigma} \left( \frac{\tau_{ij} c_j}{a^\theta P_i} \right)^{(1-\sigma)} E_i - f_{ij} \quad (3)$$

- Effect of fixed and variable trade cost : Both depends on parameter  $\theta$ 
  - Positive if  $\theta > 1$ ,

$$\frac{\partial a_{ij}^*}{\partial f_{ij}} = \frac{1}{\theta(\sigma - 1)} f_{ij}^{\frac{1}{\theta(\sigma-1)} - 1} \left[ \frac{1}{\sigma - 1} \left( \frac{\sigma}{E_i} \right)^{1/(\sigma-1)} \frac{\tau_{ij} c_j}{P_i} \right]^{1/\theta} > 0$$

$$\frac{\partial a_{ij}^*}{\partial \tau_{ij}} = \frac{1}{\theta} \tau_{ij}^{(1/\theta) - 1} \left[ \left( \frac{1}{\sigma - 1} \right) \left( \frac{\sigma f_{ij}}{E_i} \right)^{1/(\sigma-1)} (c_j / P_i) \right]^{1/\theta} > 0$$

## Comparative Statistics Results

- Productivity “a” follows pdf  $g(a)$  and cdf  $G(a)$  : assume truncated Pareto distribution  $[a_L, a_H]$
- Cut-off productivity  $a_{ij}^*$  where  $\pi_{ij}(a_{ij}^*) = 0$
- Trade volume

$$V_{ij} = \begin{cases} \int_{a_{ij}^*}^{a_H} a^{(\sigma-\theta)} dG(a), & \text{for } a_{ij}^* \leq a_H \\ 0, & \text{otherwise} \end{cases} \quad (4)$$

Then, trade value

$$M_{ij} = \left(\frac{\sigma}{\sigma-1}\right)^{1-\sigma} \left(\frac{c_j \tau_{ij}}{P_i}\right)^{1-\sigma} E_i N_j V_{ij} \quad (5)$$

- Use trade value to infer the relationship between trade costs and cutoff productivity

$$\frac{\partial M_{ij}}{\partial \tau_{ij}} = -\sigma^{1-\sigma} (\sigma-1)^\sigma \left(\frac{c_j \tau_{ij}}{P_i}\right)^{-\sigma} \tau_{ij}^{-\sigma} E_i N_j V_{ij} < 0$$

$$\frac{\partial M_{ij}}{\partial \tau_{ij}} = \frac{\partial M_{ij}}{\partial a_{ij}^*} * \frac{\partial a_{ij}^*}{\partial \tau_{ij}} < 0$$



# Empirical estimation

- Two-stage estimation (Helpman et al., 2008)
- Disaggregated product-importer-exporter level

## 1 Selection equation

$$\rho_{hij} = Pr(T_{hij} = 1) = \Phi(\xi_j^* + \xi_h^* + \gamma_1^* \ln DIST_{ij} + \gamma_2^* ADJ_{ij} + \gamma_3^* COMLANG_{ij} + \gamma_4^* \ln RTA_{ij} + \kappa_1^* Gov_i + \kappa_2^* SPS_{hij}) \quad (6)$$

## 2 Trade equation

$$\ln m_{hij} = \psi_0 + \psi_{ih} + \psi_{jh} + \gamma_1 \ln DIST_{ij} + \gamma_2 ADJ_{ij} + \gamma_3 COMLANG_{ij} + \gamma_4 \ln RTA_{ij} + \ln(\exp[\delta(\hat{z}_{hij}^* + \hat{\lambda}_{hij}^*) - 1] + \beta_{u\eta} \hat{\lambda}_{hij}^* + e_{hij}) \quad (7)$$

where  $\beta_{u\eta} \equiv \text{corr}(u_{hij}\eta_{hij})/(\sigma_u/\sigma_\eta)$

$\ln(\exp[\delta(\hat{z}_{hij}^* + \hat{\lambda}_{hij}^*) - 1])$  : correct for absence of extensive margin ( # of exporting firms through expected probability)

$\hat{\lambda}_{hij}^*$ : inverse Mills Ratio for correcting sample selection error

- Cross section data for 2012
- Food and agricultural product trade value and quantity data from FAO, trade cost data from CEPII, standards data from Worldbank and WTO
  - Exclusion restrictions should determine probability of exporting but not affect trade value
    - Governance indicators: quality of regulations, governmental efficiency, rule of law (Worldbank)
    - Sanitary and phytosanitary standards (SPS) data from WITS and I-TIP

# Estimation Results

	OLS	PPML	Probit (Mij>0)	Hetero firm NLS
In Dist	-1.225*** (0.012)	-0.736*** (0.020)	-0.358*** (0.005)	-0.976*** (0.338)
ADJ	1.014*** (0.032)	0.738*** (0.039)	0.700*** (0.019)	0.372** (0.159)
LANG	0.690*** (0.022)	0.349*** (0.042)	0.318*** (0.009)	0.493 (0.076)
RTA	0.650*** (0.023)	0.963*** (0.045)	0.166*** (0.010)	0.530*** (0.051)
SPS			-0.075*** (0.025)	
Governance			1.178*** (0.016)	
Delta				0.598*** (0.107)
Inv Mills ratio				0.088 (0.823)
Importer FE	No	No	No	No
Exporter FE	No	No	Yes	No
Product FE	No	No	Yes	No
Importer-product FE	Yes	Yes	No	Yes
Exporter-product FE	Yes	Yes	No	Yes
Observations	106,462	279,799	279,799	106,462
Wald chi2	-	-	73,895	
Adj R-squared	0.5845	-	0.3112	0.6074

Clustered robust standard errors in parentheses

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

## Estimation Results

- Estimates follow theoretical expectation
  - Trade determinants (distance, adjacency, language)
  - SPS negatively influence trade flows
- Conventional gravity model estimation would be biased upward
- By introducing non-linear coefficient  $\delta$  and inverse Mills ratio, coefficients of trade determinants become consistent

## Conclusion

- Effect of trade costs depend on “scope for quality differentiation”
  - Product quality as well as firm productivity are determinants of export threshold
  - Increasing in trade costs reduces extensive margin by increasing export threshold
- Empirical evidence supports argument that fixed costs, SPS, negatively affect probability exporting
  - Ignoring control of heterogeneity and sample selection leads to bias in estimating effect of variable trade cost

## Appendix: Selection equation

- Latent variable

$$Z_{hij} \equiv \frac{\frac{1}{\sigma} \left(\frac{\sigma}{\sigma-1}\right)^{1-\sigma} (\tau_{hij} c_{hj} a^{-\theta})^{1-\sigma} E_{hi} P_{hi}^{\sigma-1}}{f_{hij}} = \left(\frac{a_{hij}}{aL}\right)^{-\theta(1-\sigma)} \quad (8)$$

- Ratio of export profit to fixed cost
- The reduced form of selection equation

$$\ln z_{hij} = \xi_0 + \xi_{hi} + \xi_{hj} - \gamma d_{hij} - \kappa \phi_{hij} + \eta_{hij} \quad (9)$$

$$(1 - \sigma) \ln \tau_{hij} = \gamma d_{hij} + u_{hij} \quad \& \quad \ln(f_{hij}) = \vartheta_{hi} + \vartheta_{hj} + \kappa \vartheta_{hij} + v_{hij}$$
$$\eta_{hij} \sim N(0, \sigma_u^2 + \sigma_v^2)$$

- Since  $z_{hij}$  is not observed directly, we set up indicator function  $T_{hij}$

$$\rho_{hij} = \Pr(T_{hij} = 1 | \text{observed variables}) = \Pr(T_{hij} = 1 | \xi_0 + \xi_{hi} + \xi_{hj} - \gamma d_{hij} - \kappa \phi_{hij} > -(v_{hij} + u_{hij})) \quad (10)$$

$$\Phi(\xi_0^* + \xi_{hi}^* + \xi_{hj}^* - \gamma^* d_{hij} - \kappa^* \phi_{hij}) = \Phi(\hat{z}_{hij}^*) = \Phi(X_{hij} \vartheta^*)$$

\* indicates estimates divided by the standard deviation of  $(v_{hij} + u_{hij})$

## Appendix: Selection equation

$$M_{hij} = \left( \frac{C_{hj} \tau_{hij}}{P_{hi}} \right)^{1-\sigma} E_{hi} N_{hj} V_{hij}$$

$$\text{where } V_{hij} = \frac{k a_L^{k-\theta(1-\sigma)}}{k - \theta(1-\sigma)(a_H^k - a_L^k)} W_{hij}$$

$$W_{hij} \text{ indicates } \max \left\{ \left( \frac{a_{hij}}{a_L} \right)^{k-\theta(1-\sigma)} - 1, 0 \right\}$$

$$\text{Therefore } \ln m_{hij} = \psi_0 + \psi_{ih} + \psi_{jh} + \gamma \ln d_{hij} + w_{hij} + u_{hij}$$

- Since  $Z_{hij} = \frac{a_{hij}}{a_L}^{-\theta(1-\sigma)} \rightarrow W_{hij} = Z_{hij}^{k-\theta(1-\sigma)/\theta(1-\sigma)} - 1$

$$\hat{w}_{hij} = \ln(\exp[\delta(\hat{z}_{hij}^* + \hat{\lambda}_{hij}^*)] - 1) \quad \hat{\eta}_{hij}^* = \phi(\hat{z}_{hij}^*)/\Phi(\hat{z}_{hij}^*) \quad \hat{z}_{hij}^* = \hat{z}_{hij} + \hat{\eta}_{hij}^*$$

$$\delta = \sigma_{\eta}(k - \theta(1 - \sigma))/\theta(1 - \sigma)$$

- Trade equation

$$\begin{aligned} \ln m_{hij} = & \psi_0 + \psi_{ih} + \psi_{jh} + \gamma_1 \ln DIST_{ij} + \gamma_2 ADJ_{ij} + \gamma_3 COMLANG_{ij} + \\ & \gamma_4 \ln RTA_{hij} + \ln(\exp[\delta(\hat{z}_{hij}^* + \hat{\lambda}_{hij}^*)] - 1) + \beta_{u\eta} \hat{\lambda}_{hij}^* + e_{hij} \end{aligned} \quad (11)$$

## Appendix: Trade equation

- Trade equation

$$\ln m_{hij} = \psi_0 + \psi_{ih} + \psi_{jh} + \gamma_1 \ln DIST_{ij} + \gamma_2 ADJ_{ij} + \gamma_3 COMLANG_{ij} + \gamma_4 \ln RTA_{hij} + \ln(\exp[\delta(\hat{z}_{hij}^* + \hat{\lambda}_{hij}^*) - 1] + \beta_{u\eta} \hat{\lambda}_{hij}^* + e_{hij}) \quad (12)$$

- $\delta = \sigma_{\eta}(k - \theta(1 - \sigma))/\theta(1 - \sigma)$

Parameter	Value	Source
Elasticity of substitution ( $\sigma$ )	3.38 <sup>1</sup>	Bernard et al. (2003) Broda and Weinstein (2006)
Shape of parameter of the Pareto productivity distribution ( $k$ )	4	Bernard, Redding, Schott (2009)
Quality parameter ( $\theta$ )	1.335 $\bar{1}$ .420	Crino and Epifani (2010)

<sup>1</sup>Geometric mean of sigma for agricultural and food industries(SITC 001 112)



## Appendix: Robustness check

- Non parametric estimation to control joint normality assumption

	Hetero firm NLS	Indicator Variables (50 bin)	(100 bin)
In Dist	-0.976*** (0.090)	-1.044*** (0.023)	-0.972*** (0.034)
ADJ	0.372** (0.174)	0.563*** (0.046)	0.429*** (0.068)
LANG	0.493 (0.085)	0.603*** (0.028)	0.547*** (0.045)
RTA	0.530*** (0.062)	0.510*** (0.025)	0.476*** (0.046)
Delta	0.598*** (0.121)		
Inv Mills ratio	0.088 (0.224)		
Importer-product FE	Yes	Yes	Yes
Exporter-product FE	Yes	Yes	Yes
Observations	106,462	106,462	106,462
Wald chi2			
R-squared	0.6074	0.6104	0.6108

Robust standard errors in parentheses

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$