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FINANCIAL HEALTH OF U.S. FARM BUSINESSES A Region, Type, and Size Analysis

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FINANCIAL HEALTH OF U.S. FARM BUSINESSES A Region, Type, and Size Analysis

by

Allan E. Lines and Mitchell Morehart *

INTRODUCTION

The U.S. farm sector has undergone a period of substantive economic decline over the past five years. Farm real estate values, a reflection of farm business economic health, peaked in 1981 and have declined each year since. After a decade of rapid increase, farm debt has leveled off and only recently begun to decrease. The increasing number and amount of farm loans being liquidated and/or in a delinquent state during this period evidence the growing degree of financial ill-health on U.S. farms (Wilkinson; USDA, December 1985).

During recent years, numerous studies have enhanced the understanding of the "farm crisis" from state, regional and national perspectives (USDA, July 1985; Lines and Zulauf; Dobson, et.al.; Lines and Pelly; FAPRI). These studies examined the degree of financial stress relative to size, type, region, and other demographic characteristics by focusing on the immediacy of farm family financial stress using debt-to-asset ratios and/or cash balances for indicators. Few have gone beyond the use of descriptive statistics to appraise the short-run ability of a farm family to meet its cash flow committments. Given that these previous studies focused attention on the assessment and survival of the near-term cash crisis that farm families have been experiencing, they quite appropriately incorporated cash generating and conservation strategies commonly employed by farm families. This study, on the other hand, changes the focus to the intermediate and longer run

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and stresses profitability of the business with cashflow being a secondary consideration.

The objective of this study was to improve understanding of the U.S. "farm crisis" from a business management perspective by: (1) development of a comprehensive multi-dimensional indicator of farm business financial health and (2) rigorously analyzing cross-sectional data to determine the statistical association of region, farm type, farm size, and management ability with relative farm business financial health.

The remainder of this paper is organized as follows: (1) development of the conceptual framework and the ordinal measure of farm business financial health, (2) a discussion of the data base and estimation methods, (3) the model used to estimate sample parameters is presented, and (4) results are provided, interpreted, and summarized.

ORDINAL MEASURE OF FINANCIAL HEALTH

Assessment of financial health, from a business management perspective, includes some measure ot: (1) liquidity - the ability of the business to meet its short term financial obligations, (2) solvency - a measure of risk-bearing ability, and (3) profitability - an indicator of longer run survivability. Previous studies did not include the important element of profitability. Farm business financial health, as measured by cash balance position and/or debt-to-asset ratio, has been distorted for the following reasons: (1) non-farm earnings and unpaid family labor subsidized farm losses, (2) positive cash flows (i.e. no stress) were created by allowing farm families to "live off of depreciation" (i.e. not charging the business for a legitimate business

expense - depreciation), and (3) unaccounted-for changes in inventory temporarily disguised or created cash flow stress.

A multi-dimensional ordinal measure of farm business financial health was developed in an attempt to correct for these weaknesses. It was constructed, as per Table 1, using heuristic rules to establish criterion levels. Liquidity, solvency, and profitability, respectively, were assumed satisfactory (+) if: (1) the business could meet its short run cash needs (i.e. operating costs, operator labor charges, and principal payments), (2) debt-to-asset ratio did not exceed the point where lender and farmer were equally invested (.5), a commonly accepted maximum for many lenders, and (3) rate of return to assets was not less than the average long term current return to farm assets (.04) (Melichar); otherwise they were unsatisfactory (-).

Profitability was the primary determinant of the ordinal ranking assigned to each satisfactory/unsutisfactory combination of variables used to assess farm business financial health. Given the long-run orientation of this study, a favorable profit picture was viewed as most important since it indicates, ceteris paribus, an ability to correct unsatisfactory liquidity and/or solvency. Liquidity rather than solvency was chosen as the second criteria for ranking financial health. A business was perceived to be in a better financial state, ceteris paribus, if it did not have to borrow additional funds to meet cash obligations, regardless of profit condition, and thus commit itself to paying current expenses out of future income. Support for liquidity as the second ranking criteria is evident by noticing that an unfavorable solvency position for farms displaying favorable profit and liquidity did not

create liquidity problems (i.e. profits were great enough to service higher debt-to-asset ratios and yet maintain favorable cash balances).

DATA BASE AND ESTIMATION METHODS

The source of data was the 1984 Farm Costs and Returns Survey (FCRS). This survey was collected through joint efforts of the Statistical Reporting Service and the Economic Research Service, USDA. The FCRS was unique. It consisted of 23,386 personal interviews by 1600 trained enumerators and yielded 13,003 useable questionnaires. The data collected on farm expenditures, income, capital investments, values of stocks and inventories, and financial information relative to demographic and other farm characteristics has been extensively used by USDA, GAO, and the Federal Reserve in other financial analyses of the farm sector. The sample was a multi-frame, stratified survey consisting of list and area frames. Farms from the list frame were stratified by various criteria such as economic size and size of labor force; the area frame was stratified by land use type. The sampling scheme allowed for the construction of survey expansion factors which were equal to the inverse of the selection probability. The expanded number of farms covered by the FCRS totalled 1.7 million, compared with 2.2 million farms from the 1982 Census of Agriculture. The Census included .3 million farms with sales of less than \$1,000 which were excluded from the FCRS. Most undercounting of farms was for small sales classes. The survey provided a representative count for commercial farms (USDA, July 1985).

Data properties, in particular the ordinal nature of the dependent variable, limited the candidates for estimation technique to models of

qualitative choice. Alternative specifications within this class of models include the linear probability model, the linear discriminant model, the probit model, and the logit model.

Empirical applications of the linear probability model have been discouraged based on problems that arise when conventional linear regression principles are imposed on a specification which includes a non-continuous dependent variable (Pendyck and Rubinfeld; Judge, et.al.). The linear discriminant model has received considerable use in analyses of qualitative choice. It has been shown, however, that presence of discrete exogenous variables within the model violate the basic assumption of multivariate normality (Halperin, et.al.; Press and Wilson; Harrell and Lee; Efron). As a consequence, significance tests regarding coefficients of the linear discriminant model may yield misleading conclusions.

The probit and logit formulations are monotonic transformations which insure that predicted values of the dependent variable are confined to its range. The probit model is based on the standard normal density function while the logit model relies on the logistic density function. It is often argued that the logit model has computational advantages since it is a closed functional form with convenient curvature properties for numerical optimization. The probit model, on the other hand, has as its argument the limit of an integral which cannot be expressed in closed form. This justification for the choice of the logit model over the probit model is tenuous since with maximum likelihood estimation (as occurs in most applications) computational difficulties are virtually undistinguishable regardless of standard distribution function selected. Applications of both models may be found in recent agricultural economics

studies of qualitative choice (Chambers and Foster, probit; Garcia et.al., Lines and Zulauf, logit; and Capps and Kramer, logit and probit). Empirically, the choice between probit and logit is based on convenience, especially with respect to available computer software. In this study the logit model was applied since available software easily handled the complexities of estimation given the eight class ordinal variable.

ECONOMETRIC MODEL

LOGIST, a SAS procedure, was used to estimate the probability of a business being in financial health category j or above (Harrel, 1983; Harrel, 1985). Because of the biased nature of the sample, a weighted LOGIST procedure was used (DuMouchel, et. al.). The weights in this case correspond to the survey expansion factors described above. The weighted ordinal logistic model for this dependent variable having values 0,1,...,7, can be stated as follows for $l \leq j \leq 7$:

 $= (\alpha_j + \beta_1 X_{11} W_1 + \dots + \beta_n X_{1n} W_1)$ Probability $(Y_1 \ge j) = 1/e$

Parameters were estimated using the following log linear transformation:

logit Probability $(Y_1 \ge j) = \alpha_j + \beta_1 X_{11} W_1 + \dots + \beta_n X_{1n} W_1$ where: Y_1 = ordinal indicator for ith observation j = value of ordinal indicator α_j = intercept term X_1 = predictors for ith observation β = regression coefficients W_1 = weight variable for ith observation

A separate parameter α_j is required for each level of $Y_i = 1, 2, ..., 7$ and Probability ($Y_i = 0$) is obtained from 1-Probability ($Y_i \ge 1$). The model uses the ordering of Ys but no assumptions are made regarding the spacing of scale intervals. Other model assumptions include independence

of the Xs and linearity in each X (Harrel and Lee). The independent variables include: (1) a dummy variable for each of ten U.S. agricultural regions, as defined by the Statistical Reporting Service, (2) a dummy variable for each of ten farm types, as defined by Standard Industrial Code classifications, (3) a dummy variable for type of business organization (proprietorship or not), (4) a dummy variable for degree of enterprise specialization, (5) log of gross income (a measure of size), and (6) percent of land operated that is rented.

Corn Belt, Lake State and Northern Plains farms and cash grain, livestock, and dairy farms were hypothesized to have a higher probability of being in worse financial health (USDA, 1985). Large farms were expected to be in better financial health than smaller farms (USDA, 1985), as were businesses that rented a higher proportion of the operated acreage (Lines and Zulauf). Specialized farms and proprietorships were anticipated to be in better financial health because of the ability to spread overhead costs and the potential for closer cost control and better management, respectively.

RESULTS

Results are presented in frequency distributions, graphical analyses, and statistical summaries. Each adds a unique dimension to understanding and interpretation. Tables 2 thru 4 provide the relative incidence of financial health categories with respect to farm size, type, and region. The distributions have a bi-modal character. Farm businesses, for the most part, exhibited either quite good or quite poor financial health. The reader is reminded that this analysis excludes

off-farm earnings, but includes estimated inventory changes, depreciation allowances, and charges for family labor.

Information in Table 2 shows that (1) the financial health of commercial farms was better than for all farms and (2) the financial health of larger commercial farms was better than for smaller commercial farms. Restricting the analysis to commercial farms (at least \$40,000 gross income) resulted in a significant change in the distribution of financial health, relative to that for all farms. Whereas, nearly seventy percent of all farms had poor financial health (categories 6 and 7), only forty percent of commercial farms were so classified. Conversely, approximately forty percent of commercial farms and only twenty percent of all farms were in good financial health (categories 0 and 1). Restricting size of commercial farms even further, to a minimum of \$100,000 and then \$250,000, resulted in further shifting of the distribution away from poor financial health categories toward good. Approximately fifty-five percent of the largest commercial farms (at least \$250,000 gross income) were in good financial health while one-fourth remained in poor condition. These results suggest a strong continuum of improved financial health associated with farm size.

Results in Tables 3 and 4 are not exhaustive but clearly demonstrate (1) apparent regional and farm type differences in farm business financial health and (2) commercial farms had better financial health, irrespective of region or type of farm. On an "all farms" basis, these data indicate a worsening of farm business financial health as region changed from Northern Plains to Corn Belt to Northeast. When restricted to commercial farms, there was no apparent difference between Corn Belt

and Northern Plains farms but Northeast farms remained in worse financial condition. Commercial farms had better financial health than "all farms", regardless of the region. Similarly, on an "all farms" basis, livestock and nursery/greenhouse farms were in poorer financial health than were grain farms. However, when only commercial farms were considered, nursery/greenhouse farms were in better financial health than either grain or livestock farms and livestock farms clearly remained in the worst financial condition. Regardless of type, commercial farms exhibited better financial health than did "all farms."

The estimated coefficients in Table 5 must be interpreted with care. The left side of the logistic model is a logarithm of the odds of being classified in category j or higher, not an actual probability. However, since a logarithm is a monotonic transformation, the log odds reveal characteristics of the underlying probability (Pendyck and Rubinfeld). For this analysis it is sufficient to know the significance and relative directional change of estimated parameters to infer changes in probabilities. Coefficients of dummy variables in the analysis (region, type, specialization, and business type) act as intercept shifters of the log probability functions of a farm being in category j or greater; negative significant coefficients imply improved financial health and positive coefficients imply the opposite, relative to farms in the omitted class. Coefficients of continuous variables (gross income and percent of land rented), on the other hand, act as slope indicators of the log probability functions; negative significant coefficients, in this instance, imply improved financial health as the independent variable increases and positive coefficients imply the

opposite. Differences in the magnitude of significant coefficients indicate differential intercept shifts and slopes in the log probability functions and infer a greater or lesser change in financial health.

In the context of "all farms", farms located in the Northeast, Lake States, Southern Plains, Mountain, and Pacific regions exhibited significantly worse financial health (i.e. a higher probability of being in category j or greater) than farms in the Corn Belt ($\alpha = .01$). The financial health of farms in the remaining regions was not significantly different from that of Corn Belt farms. No region, when considering "all farms", had significantly better farm business financial health than the Corn Belt (i.e. farms in the Corn Belt were not worse off than other regions). Changing the context of the analysis to "commercial farms" resulted in the Lake States and Southern Plains regions no longer being significantly different from the Corn Belt region. Commercial farms in the Pacific region (coefficient .55) likely had a higher probability of being in category j or greater (worse financial health) than farms in the Northeast region (coefficient .39). These results did not support the hypothesis associating poorer financial health with farms in the Corn belt, Lake States, and Northern Plains regions.

On an "all farm" basis, results indicate that livestock, dairy, and other livestock farms had significantly worse financial health (a positive coefficient) than did grain farms, the omitted category. Opposite to what occurred in the regional analysis, the number of significantly different types of farming increased as minimum size was restricted to \$40,000 gross income or more. Added to the list was

nursery/greenhouse farms, that had significantly better financial health (i.e. negative coefficient) than did grain farms. These results support the hypothesis that livestock and dairy farms were likely in worse financial condition. However, grain farms were not found to be worse off than other types of farms, except as noted above.

The remaining dummy variables, specialization and business type, were significant with the correct sign, in the context of all farms, but neither was significant when only commercial farms were considered. The significant coefficients (all farms) mean that (1) farms with some but not excessive diversification and (2) those farms organized as proprietorships had better financial health, relative to completely specialized or very diversified farms and those organized as non-proprietorships. The lack of significance of these variables on commercial farms is not surprising. Smaller (non-commercial) farms would be expected to exhibit a greater degree of diversification and a higher incidence of proprietorship organization. Removal of these farms (i.e. commerical farm analysis) results in a more homogeneous sample, relative to these variables, hence the disappearance of significance.

Coefficients for gross income and percent of land rented were significant and had the proper signs. The results support the hypotheses that larger farm businesses and those that rented a greater percentage of land operated were in better financial health. These relationships, characterized by negative coefficients, are illustrated in Figures 1 thru 3. Small farm businesses had a very high (low) probability of being in financial health category 6 or greater (l or less). As farm size increased, the probability of being in financial health category 6

or greater diminished from near 1 to .3; conversely, the probability of being in category 1 or less increased from near 0 to .5. Similar relationships existed for percent of land rented in the contexts of all and commercial farms.

Model validity was substantiated by the likelihood ratio chi-squares of 1474 and 505. The predictive ability of the model is assessed by examining its rank correlation statistic that has a range of 0 (no predictive ability) to 1 (perfect predictive ability). The statistics for this model were .71 and .63, not unreasonable for cross-sectional analysis.

SUMMARY AND IMPLICATIONS

These findings indicate that the "farm crisis", when examined from a comprehensive view of business financial health, is more severe than other literature would suggest. Using the same data as this analysis but by only addressing the near-term cash crisis USDA concluded that twelve percent of all farms were financially stressed (USDA, July 1985). The Federal Reserve, again using the same data base, suggested that seventeen percent of the commercial farms were financially stressed (Melicher, October 1985). Nearly seventy percent of all U.S. farm business and forty percent of commercial farms were in serious financial difficulty, when oft-farm income, inventory changes, depreciation, and unpaid family labor were taken into account. This is not to say that these percentages will fail in the near or not too distant future. Many farm businesses, commercial included, will purposely continue to be subsidized by off-farm income and unpaid family labor and many will survive for some time by delaying the replacement of depreciable assets.

Relative to other farm types, financial health was poorest on dairy and livestock farms. Surprisingly, however, grain farms and farms located in the Corn Belt or Northern Plains regions were not worse off than most others; however, when restricted to commercial farms, nursery/greenhouse farms were significantly better. Increased farm size and a higher portion of land rented significantly increased the probability of having good financial health. A limited degree of diversification and being organized as a proprietorship were positively associated with financial health only on an "all farms" basis. Neither variable was significant on commercial farms.

The results have important implications for agricultural policy. Poor farm business financial health is pervasive in U.S. agriculture, although clearly worse in some regions, on some types of farms, and on smaller farms. Tightly targeted economic assistance will (1) only address part of the problem and (2) result in proliferation of costly specific programs that may reward poor and/or part-time managers that may not warrant, need, or desire assistance. A broad spectrum approach designed to shift the distributions in Tables 2 thru 4 upwards will (1) have high unacceptable public cost, (2) encourage over investment in agriculture, and (3) result in over production, low incomes, and poor financial health. Policies to assist operators of smaller farm businesses inadvertently subsized by off-farm income, unpaid family labor, and/or asset depletion to understand broad economic issues and problems and adjust to economic realities may be useful. Likewise, policies to encourage development, implementation, and participation of farmers in educational and assistance programs that emphasize understanding,

attainment, and maintenance of good farm business financial health may be beneficial.

Much of the financial ill-health identified in this analysis results from (1) exclusion of off-farm income and (2) inclusion of depreciation. Both are important from a policy perspective. Policies grounded in the concept that the economic well-being of farm businesses includes off-farm income, foster a farm sector dependent upon off-farm income and unable to pay all its expenses. Such policies inadvertently promote transition to a sector dominated by part-time and commercial farms that are subsidized by family members working outside the sector. From a policy perspective the farm operators themselves, rather than the general public, absorb the hidden costs of ensuring an adequate food supply. It remains to be determined if this is desirable or equitable as a policy objective. On the other hand, policies dependent upon a definition and measurement of farm financial stress that excludes depreciation underestimate the severity of the crisis, exacerbate the transition to part-time farms, and jeopardize sector productivity. Currently and during the past five years many farm businesses have avoided being classified as "financially stressed" by not replacing depreciable assets consumed in production. The sector ability to meet cash flow needs via this strategy is rapidly disappearing. If current conditions persist many commerical operators will exhaust their ability to "live off depreciation" and find themselves unable to continue in business. Policies designed to address the "farm financial crisis" without considering these components of the problem may be short-sighted and unable to deal with continued economic deterioration in the farm sector.

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The weighted ordinal logistic regression methodology and model developed and used in this study provides a powerful analytical tool for researchers and policy analysts concerned about the financial well-being of the U.S. farm sector. Using the model it is possible to predict, with seventy percent accuracy, the probability of a farm having good, fair, or poor financial health, given a distinct set of exogenous variables. Further improvement in the data base and inclusion of other exogenous variables in the model, including indicators of management ability, will likely improve model performance. Data limitations precluded inclusion of such variables in this analysis. The model provides a mechanism for synthesizing global and/or restricted sensitivity analyses with respect to (1) assumptions in the model, (2) satisfactory/unsatisfactory criterion levels for variables used to assess farm business financial health, and (3) the impact of policy variables in an attempt to determine how each would affect the distribution of farms in alternative states of financial health. Continued use of this econometric technique will permit the investigation of a critical policy question -- "Is the current farm financial situation a temporary crisis or a new norm?" The answer to this question will play a vital role in determining the development of tuture policy objectives and programs.

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leasure	Liquidity 	Solvency	Profitability	Level of Ordinal
Variable	Adjusted Cash Balance	Debt-to-Asset Ratio	Rate of Return to assets	Indicator Y _i
Criterion	$\frac{> 0 (+)}{< 0 (-)}$	<u>< .5 (+)</u> <u>> .5 (-)</u>	> 4% (+) < 4% (-)	
	+	+	+	0 (best)
Alternative Combinations of Measures	+	• ·	+	1
		+	+	2
	-	-	+	3
	+	+	-	4
	+			5
	-	+	•	6
	*	-	•	7 (worst)

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Table 1.

FARM BUSINESS FINANCIAL HEALTH

+ Satisfactory - Unsatisfactory

Table 2.

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FARM BUSINESS FINANCIAL HEALTH SIZE COMPARISON U.S. FARMS - JANUARY 1985

Financial Health	A11	Gross Income At Least		
Category	Farms	\$40,000	\$100,000	\$250,000
		Percent of	farms <u>l</u> /	
0 (best)	15	31	35	42
0 (best)	4	7	10	13
	1	3	4	5
2	2	4	5	6
3	10	14	10	6
4	*	*	*	*
5	59	29	25	19
6 7 (worst)	9	11	11	9

* Less than 1 percent

1/ Percentages may not add to 100 due to rounding

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FARM BUSINESS FINANCIAL HEALTH REGIONAL COMPARISON U.S. FARMS - JANUARY 1985

tie versient)		All Farms		Col	Commercial Farms 1/	1/
Health Lategory	Corn Belt	Northeast	N. Plains	Corn Belt	Northeast	N. Plains
			Perc	Percent of farms $2/$.	1 1 1	
	ŝ	01	27	33	31	35
U (Dest)	2, r		¢	14	5	30
ļ	~ ~	4 4	• •		*	(4
•4	7		. ~	ę		6
<i></i>	0 C	-ı at) []	10	10	13
Ŧ 1	2 +) #	1 - 17	1	*	ŧ
ŝ	- 07	11	40	23	47	23
6 7 (worst)	11	. 80	11	10	8	13

* Less than 1 percent 1/ At least \$40,000 gross income 2/ Percentages may not add to 100 due to rounding 121 × . -

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FARM BUSINESS FINANCIAL HEALTH FARM TYPE COMPARISON U.S. FARMS - JANUARY 1985

Table 4.

Nurserv	71001NN		44 20 18 128 128
Commercial Farms 1/	LIVESTOCK	1 1 1 1	24 5 11 12 12
1	Grain	Percent of farms $2/$	40 9 3 3 2 4 2 4
	Nursery	• • •	L1 11 11 11 11
All Farms	Livestock		82448*26
	Grain		27 2 4 11 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2
	rinanciai Health Category		U (best) 1 2 3 4 4 5 6 7 (worst)

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* Less than 1 percent 1/ At least \$40,000 gross income 2/ Percentages may not add to 100 due to rounding

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Table 5.

ESTIMATED CHANGES IN LOG PROBABILITIES FINANCIAL HEALTH CATEGORIES U.S. FARMS - JANUARY 1985

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	A11	Farms	Commercial Farms		
Variable	Beta	Chi-Square	Beta	Chi-Square	
Region 1/					
Northeast	.39*	10.8	.39*	3.8	
Lake States	.57*	36.2	. 14	1.8	
Northern Plains	.14	2.2	06	.4	
Appa lachia	.06	.4	24	3.1	
Southeast	.15	1.4	.19	1.4	
Delta States	.18	2.1	. 14	.7	
Southern Plains	.27*	7.9	.17	1.6	
Mountain	.67*	29.2	.45*	13.1	
Pacific	.49*	13.1	.55*	10.7	
Type of Farm <u>2</u> /					
Field Crops	.16	2.1	09	.4	
Vegetable/Melon	.19	1.0	03	.0	
Fruit/Tree Nut	.17	1.0	.14	.3	
Nursery/Greenhouse	.18	.9	-1.26*	19.7	
General Crop	. 13	2.1	. 26	2.6	
General Lives tock	.60*	71.0	.61*	52.9	
Dairy	.75*	58.9	.65*	47.5	
Poultry/Egg	.11	.3	46	5.9	
Other Livestock	1.05*	25.4	1.51*	17.7	
Gross Income	53*	738.2	57*	162.0	
Percent of Land Rented	29*	16.1	50*	32.0	
Specialization	27*	23.7	14	4.9	
Business Type	. 25*	9.5	. 10	1.3	
Model Statistics					
Chi-Square with 22 d.f.		1474	505		
p Value		.0000		.0000	
Rank Correlation Index		.71	.63		

Significant at $\alpha = .01$ *

 $\frac{1}{2}$ Omitted Region : Corn Belt $\frac{2}{2}$ Omitted Type : Grain

FINANCIAL HEALTH VS GROSS INCOME ALL U.S. FARMS - JANUARY 1985

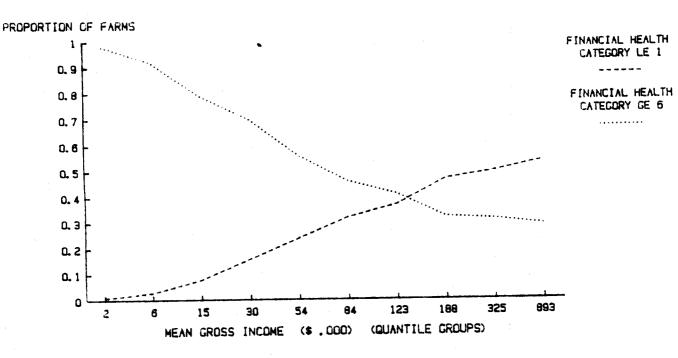


Figure 1.

Figure '.

FINANCIAL HEALTH VS PERCENT LAND RENTED ALL U.S. FARMS - JANUARY 1985

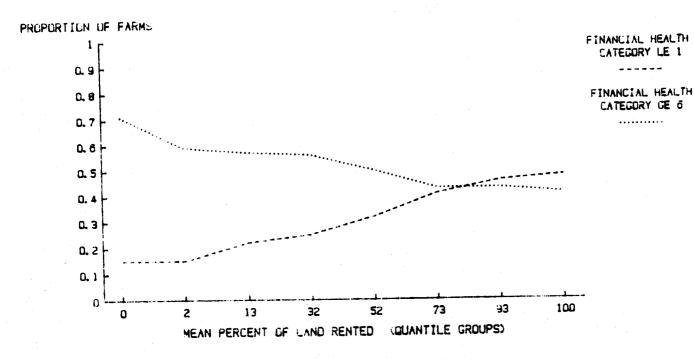
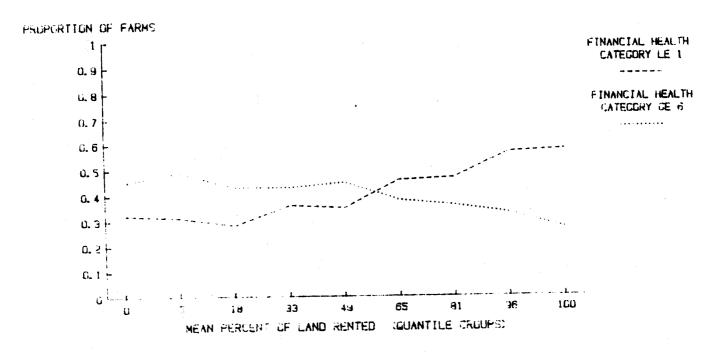


Figure 3.

FINANCIAL HEALTH VS PERCENT LAND RENTED COMMERCIAL U.S. FARMS - JANUARY 1985



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